

Homework 3 - Solutions

1.1

- a) MA(1)
AR(1)
RW with drift

b)

ACF for MA(1): Non-zero for first autocorrelation, zero for 2nd and beyond autocorrelations

$$\text{ACF}(1) = 0.5$$

$$\text{ACF}(k) = 0 \quad \text{for } k > 1$$

ACF for AR(1): Tail-off

$$\text{ACF}(1) = 0.8$$

$$\text{ACF}(2) = 0.8^2$$

$$\text{ACF}(3) = 0.8^3$$

ACF for RW: Constant and equal to 1.

$$\text{ACF}(k) = 1 \quad \text{for } k = 1, 2, 3, \dots$$

- c) RW with drift. Unpredictable returns
Predicting the stock market best: The AR(1) model.
- d) MA(1): Impact of shock lasts only one period.
AR(1): Impact of shock last a long time, but impact decays at rate 0.8.
RW: Impact of shock last forever.

1.2

a. To check for stationarity, we look at the roots of $\phi(z)$:

The roots of $\phi(z) = 1 - .803z - .682z^2$ should lie outside the unit circle.

Let's use R:

```
x1 <- c(1, -.803, -.682) # define the coefficients of polynomial
```

```
> polyroot(x1)
```

```
[1] 0.7577133+0i -1.9351326-0i
```

Conclusion: We have two roots: $z_1 = 0.7577133$; $z_2 = -1.9351326$. One is inside (z_1), one is outside (z_2) \Rightarrow non-stationary process

b. Constant moments. The distribution does not change with time.

c. Patterns

ACF: Tailing-off at a low rate.

PACF: Cut-off after lag 2

d. Produce one-, two-, and three-step ahead forecasts.

$$\hat{Y}_{t+1} = .803 Y_t + .682 Y_{t-1}$$

$$\hat{Y}_{t+2} = .803 \hat{Y}_{t+1} + .682 Y_t$$

$$\hat{Y}_{t+3} = .803 \hat{Y}_{t+2} + .682 \hat{Y}_{t+1}$$

1.3

a. $SE = 1/\sqrt{500} = \mathbf{0.04472136} \Rightarrow$ Significant ACF: 1, 5
Significant PACF: 1, 2, 3

b. $LB(5) = 500 * (500+2) * (0.307^2/(500-1) + 0.013^2/(500-2) + 0.086^2/(500-3) + 0.031^2/(500-4) + 0.197^2/(500-5)) = \mathbf{71.39342} \Rightarrow$ very significant at 5% level. That is, we reject H_0 : all five autocorrelations are jointly zero.

c. PACFs are tailing off, MA process likely. With the information above, possible MA(5) process.

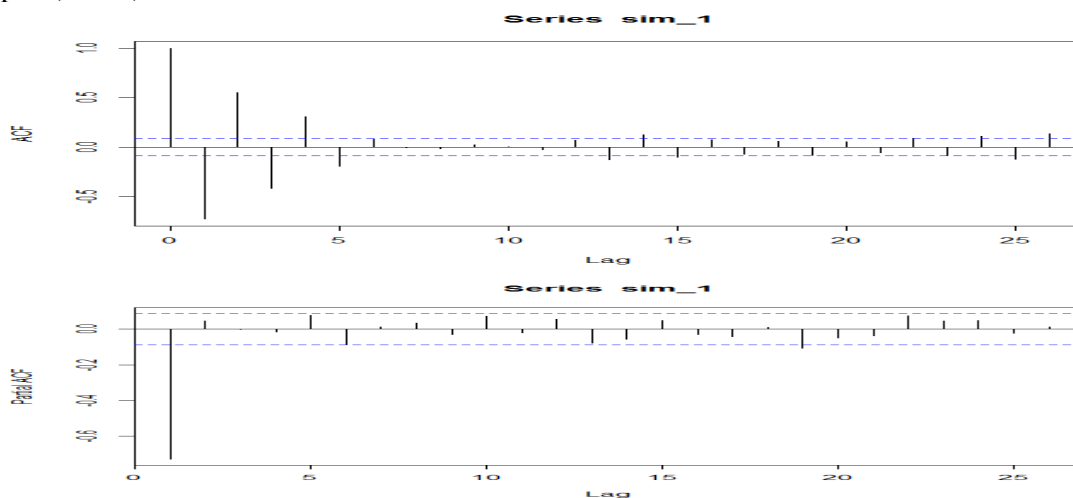
1.4

a. Plot the ACF/PACF for each series.

```
SIM_da <- read.csv("http://www.bauer.uh.edu/rsusmel/4397/Sim.arima.csv", head=TRUE,
sep=",")
sim_1 <-SIM_da$sim_1
sim_2 <-SIM_da$sim_2
sim_3 <-SIM_da$sim_3
sim_4 <-SIM_da$sim_4
```

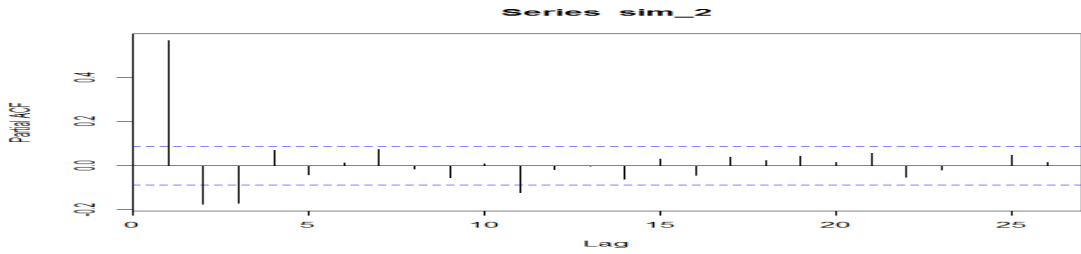
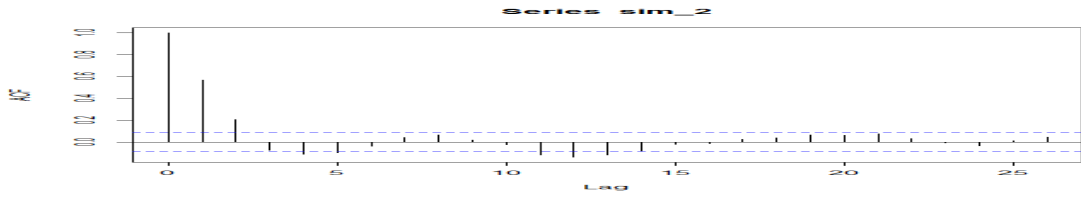
Sim_1

```
acf(sim_1)
pacf(sim_1)
```



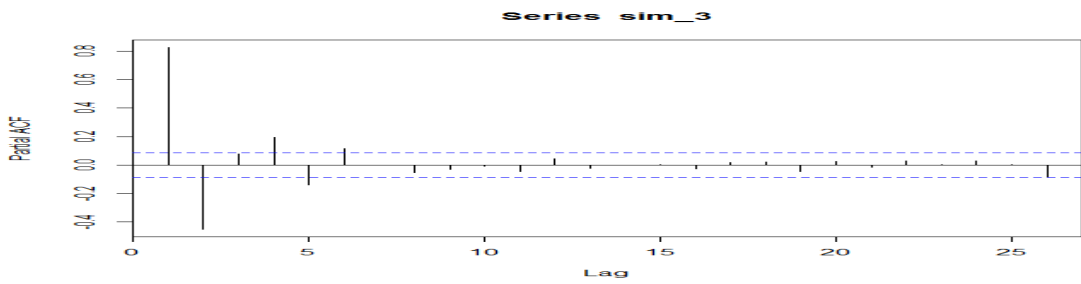
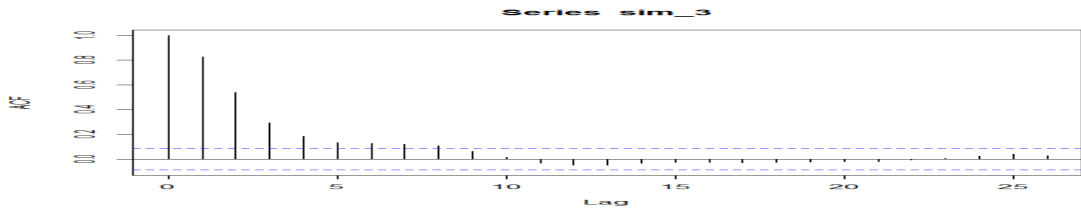
Sim_2

```
acf(sim_2)
pacf(sim_2)
```



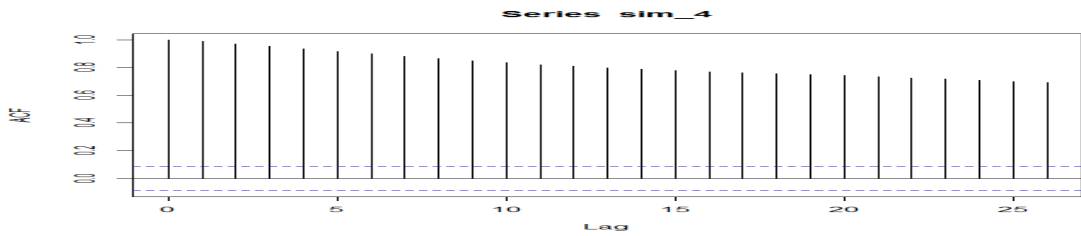
Sim_3

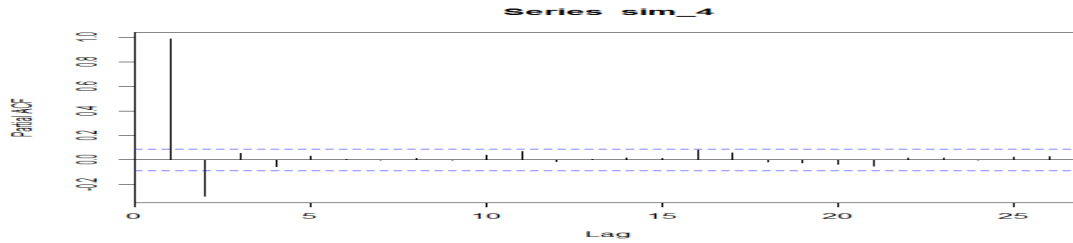
acf(sim_3)
pacf(sim_3)



Sim_4

acf(sim_4)
pacf(sim_4)





b. Using the ACF/PACF guess the order of the ARIMA process

Sim_1: AR(1)

Sim_2: MA(2)

Sim_3: ARMA(1,2)

Sim_4: ARIMA(0,1,1)

c. Check the residuals to see if there is any autocorrelation left in each of your series.

Sim_1

```
fit_1 <- arima(sim_1, order=c(1,0,0))
```

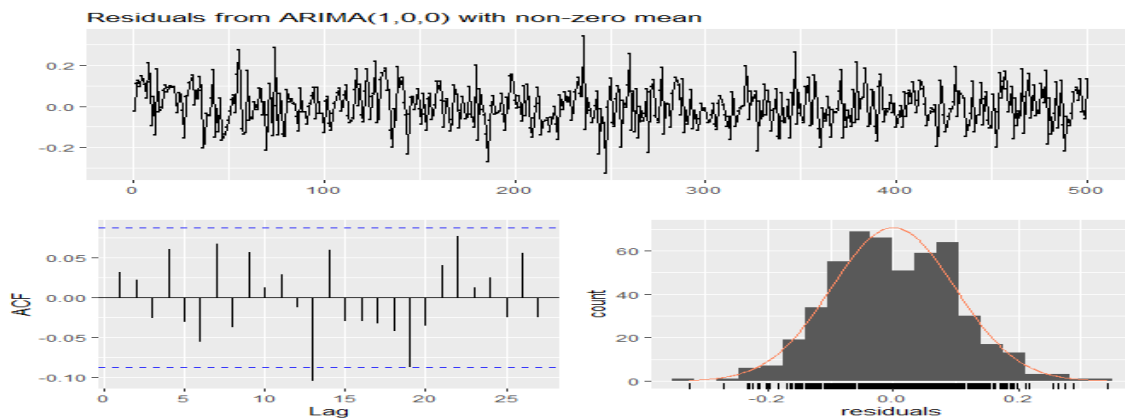
```
> fit_1
```

Coefficients:

```
ar1 intercept
-0.7307 -0.0007
s.e. 0.0305 0.0026
```

sigma^2 estimated as 0.009822: log likelihood = 445.93, aic = -885.85

```
> checkresiduals(fit_1)
```



Repeat for the other series.

1.5 (Forecasting with ARMA and SES)

a. Forecasts. We use the formula:

$$\hat{Y}_{T+l} = 0.036 + .69 \hat{Y}_{T+l-1} + 0.42 E[\varepsilon_{T+l-1}|IT] + E[\varepsilon_{T+l-q}|IT]$$

$$\hat{Y}_T = 0.036 + .69 * 3.4 + 0.42 * (-1.3) = 1.836$$

$$\hat{Y}_{T+1} = 0.036 + .69 * 1.836 = 1.30284$$

$$\hat{Y}_{T+2} = 0.036 + .69 * 1.30284 = 0.9349596$$

$$b. \text{MSE}_{\text{arma}} = ((1.836 - (-0.032))^2 + (1.30284 - 0.961)^2 + (0.9349596 - 0.203)^2)/3 =$$

$$= \mathbf{1.380681}$$

$$c. S_t = S_{t-1} + .15 * (Y_{t-1} - S_{t-1})$$

$$\hat{Y}_T = S_T = 0.0305 + .15 * (3.4 - 0.0305) = 0.535925$$

$$\hat{Y}_{T+1} = 0.535925$$

$$\hat{Y}_{T+2} = 0.535925$$

$$\text{MSE}_{\text{ses}} = ((0.535925 - (-0.032))^2 + (0.535925 - 0.961)^2 + (0.535925 - 0.203)^2)/3 =$$

$$= \mathbf{0.2046889}$$

d. SES has lower MSE => Preferred model.

1.6 (Forecasting Evaluation)

a.

$$A: \hat{x}_{T+1} = 0.38 + .10 * (-0.02) = 0.378$$

$$\hat{x}_{T+2} = 0.38$$

$$\hat{x}_{T+3} = 0.38$$

$$\hat{x}_{T+4} = 0.38$$

$$\hat{Y}_{T+l} = 0.38 + .10 \hat{Y}_{T+l-1} - 0.09 \hat{Y}_{T+l-2}$$

$$B: \hat{x}_{T+1} = 0.63 + .17 * 0.31 - .009 * 0.02 = 0.68252$$

$$\hat{x}_{T+2} = 0.63 + .17 * 0.68252 - .009 * 0.31 = 0.7432384$$

$$\hat{x}_{T+3} = 0.63 + .17 * 0.7432384 - .009 * 0.68252 = 0.7502078$$

$$\hat{x}_{T+4} = 0.63 + .17 * 0.7502078 - .009 * 0.7432384 = 0.7508462$$

b.

$$\text{MAE}_{\text{ma1}} = (|0.378 - (0.62)| + |0.38 - (0.19)| + |0.38 - (-0.32)| + |0.38 - (0.72)|)/4$$

$$= \mathbf{0.368}$$

$$\text{MAE}_{\text{ar2}} = (|0.68252 - (0.62)| + |0.7432384 - 0.19| + |0.7502078 - (-0.32)| + |0.7508462 - (0.72)|)/4 = \mathbf{0.4292}$$

Conclusion: MA model has better performance.

1.7

a.

```
Sh_da <- read.csv("C:/Financial Econometrics/Real_Estate_2019.csv", head=TRUE, sep=",")
```

```
x_dat <- Sh_da$DATE
x_sf1 <- Sh_da$SF_c
x_sf <- x_lv[1:322]
```

```
fit_ar3 <- arima(x_sf, order=c(3,0,0))
> fit_ar3
```

Call:

```
arima(x = x_sf, order = c(3, 0, 0))
```

Coefficients:

	ar1	ar2	ar3	intercept
	1.0141	0.0508	-0.1791	0.0029
s.e.	0.0546	0.0787	0.0546	0.0022

sigma^2 estimated as 2.175e-05: log likelihood = 1270.64, aic = -2531.29

b.

```
T_sf <- length(res_sf_ar3)
Feb <- rep(c(1,0,0,0,0,0,0,0,0,0,0,0),T_lv/12+1) # Create January dummy
Mar <- rep(c(0,1,0,0,0,0,0,0,0,0,0,0),T_lv/12+1) # Create February dummy
Apr <- rep(c(0,0,1,0,0,0,0,0,0,0,0,0),T_lv/12+1)
May <- rep(c(0,0,0,1,0,0,0,0,0,0,0,0),T_lv/12+1)
Jun <- rep(c(0,0,0,0,1,0,0,0,0,0,0,0),T_lv/12+1)
Jul <- rep(c(0,0,0,0,0,1,0,0,0,0,0,0),T_lv/12+1)
Aug <- rep(c(0,0,0,0,0,0,1,0,0,0,0,0),T_lv/12+1)
Sep <- rep(c(0,0,0,0,0,0,0,1,0,0,0,0),T_lv/12+1)
Oct <- rep(c(0,0,0,0,0,0,0,0,1,0,0,0),T_lv/12+1)
Nov <- rep(c(0,0,0,0,0,0,0,0,0,1,0,0),T_lv/12+1)
Dec <- rep(c(0,0,0,0,0,0,0,0,0,0,1,0),T_lv/12+1)
seas_sf <- cbind(Feb, Mar, Apr, May, Jun, Jul, Aug, Sep, Oct, Nov, Dec)
seas <- seas_sf[1:(T_sf),]
fit_sf_seas <- lm(res_sf_ar3 ~ seas)
> summary(fit_sf_seas)
```

Call:

```
lm(formula = res_sf_ar3 ~ seas)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.0154680	-0.0025181	0.0000796	0.0026429	0.0118629

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.0009787	0.0008446	1.159	0.24744
seasFeb	-0.0011926	0.0011834	-1.008	0.31433
seasMar	0.0025801	0.0011834	2.180	0.02999 *
seasApr	0.0030514	0.0011834	2.579	0.01038 *
seasMay	-0.0002333	0.0011834	-0.197	0.84387
seasJun	-0.0009945	0.0011834	-0.840	0.40136
seasJul	-0.0029739	0.0011834	-2.513	0.01248 *
seasAug	-0.0020289	0.0011834	-1.714	0.08745 .
seasSep	-0.0035327	0.0011834	-2.985	0.00306 **
seasOct	-0.0024083	0.0011834	-2.035	0.04269 *
seasNov	-0.0024161	0.0011834	-2.042	0.04203 *

```
seasDec -0.0014500 0.0011945 -1.214 0.22570
```

```
---
```

```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 0.004307 on 310 degrees of freedom
```

```
Multiple R-squared: 0.1788, Adjusted R-squared: 0.1496
```

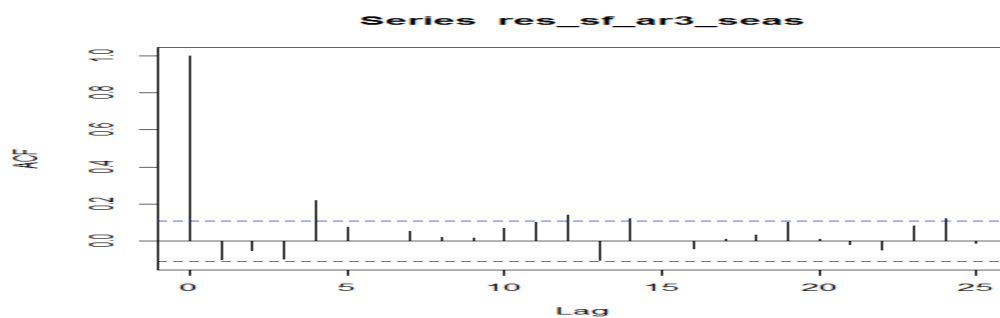
```
F-statistic: 6.135 on 11 and 310 DF, p-value: 4.064e-09
```

Conclusion: Yes, strong evidence for seasonality. A lot of the dummies are significant (the ones in red). The F-stat=6.135 rejects the H_0 that all seasonal coefficients are 0.

c.

```
res_sf_ar3_seas <- fit_sf_seas$residuals
```

```
> acf(res_sf_ar3_seas)
```



```
library(forecast)
```

```
fit_auto <- auto.arima(res_lv_ar3_seas)
```

```
> fit_auto
```

```
Series: res_sf_ar3_seas
```

```
ARIMA(2,0,2) with zero mean
```

```
Coefficients:
```

```
ar1 ar2 ma1 ma2
```

```
0.9903 -0.8376 -1.1248 0.8890
```

```
s.e. 0.0631 0.0828 0.0530 0.0646
```

```
sigma^2 estimated as 1.699e-05: log likelihood=1313.16
```

```
AIC=-2616.31 AICc=-2616.12 BIC=-2597.44
```

```
> checkresiduals(fit_auto)
```

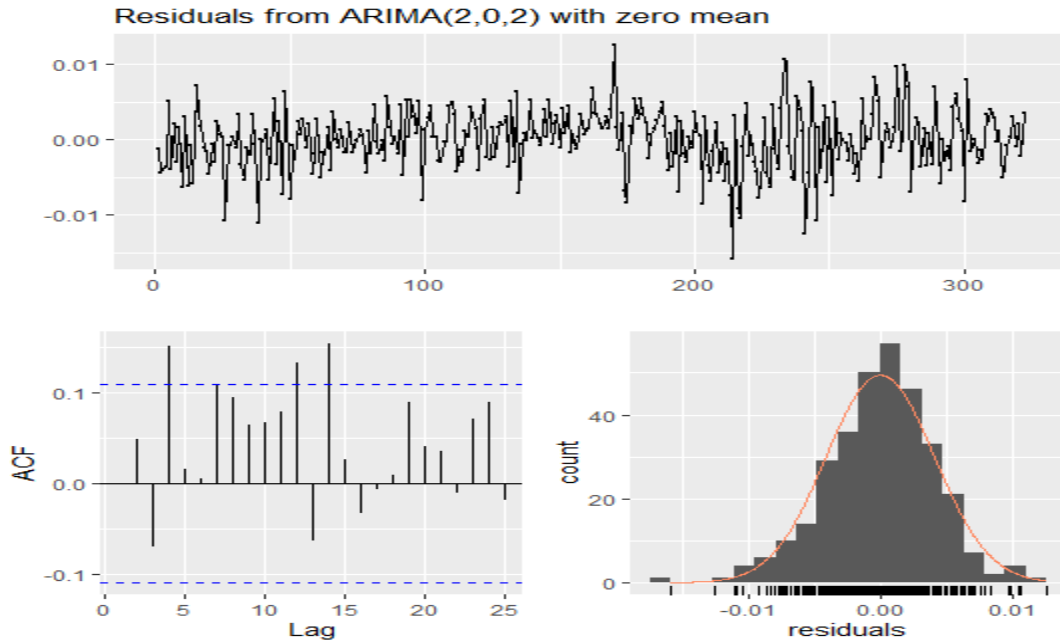
Ljung-Box test

```
data: Residuals from ARIMA(2,0,2) with zero mean
```

```
Q* = 19.659, df = 6, p-value = 0.003184
```

```
Model df: 4. Total lags used: 10
```

```
> checkresiduals(fit_auto)
```



Conclusion: Better ACF, but still with significant autocorrelations.

I try an AR(12):

```
fit_ar12 <- arima(res_sf_ar3_seas, order=c(12,0,0))
> fit_ar12
```

Call:

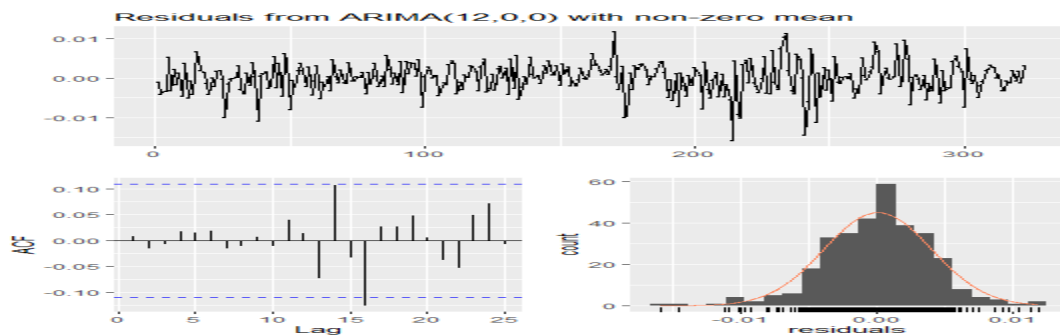
```
arima(x = res_sf_ar3_seas, order = c(12, 0, 0))
```

Coefficients:

	ar1	ar2	ar3	ar4	ar5	ar6	ar7
	-0.1458	-0.0848	-0.1135	0.1975	0.0979	0.0217	0.0820
s.e.	0.0549	0.0551	0.0551	0.0554	0.0565	0.0564	0.0565
	ar8	ar9	ar10	ar11	ar12	intercept	
	-0.0042	0.0277	0.0988	0.1136	0.1677	0e+00	
s.e.	0.0564	0.0554	0.0551	0.0551	0.0550	4e-04	

sigma² estimated as 1.553e-05: log likelihood = 1325.36, aic = -2622.71

```
> checkresiduals(fit_ar12)
```



d.

```
f_ar3_coeff <- fit_ar3$coef
T <- length(x_sf1)
y <- x_sf1[4:T]
sf_1 <- x_sf1[3:(T-1)]
sf_2 <- x_sf1[2:(T-2)]
sf_3 <- x_sf1[1:(T-3)]

x0 <- matrix(1,(T-3),1)
xx_sf <- cbind(sf_1, sf_2, sf_3, x0)
y_f0 <- xx_sf[323:nrow(xx_sf),]%% sf_ar3_coeff
err_f0_ar3 <- y[323:nrow(xx_sf)] - y_f0
mse_ar3 <- sum(err_f0_ar3^2)/length(y_f0)
> mse_ar3
[1] 5.497108e-05
```

1.8

a.

```
library(forecast)
fit_auto_8a <- auto.arima(x_sf)
> fit_auto_8a
Series: x_sf
ARIMA(4,0,2) with zero mean
```

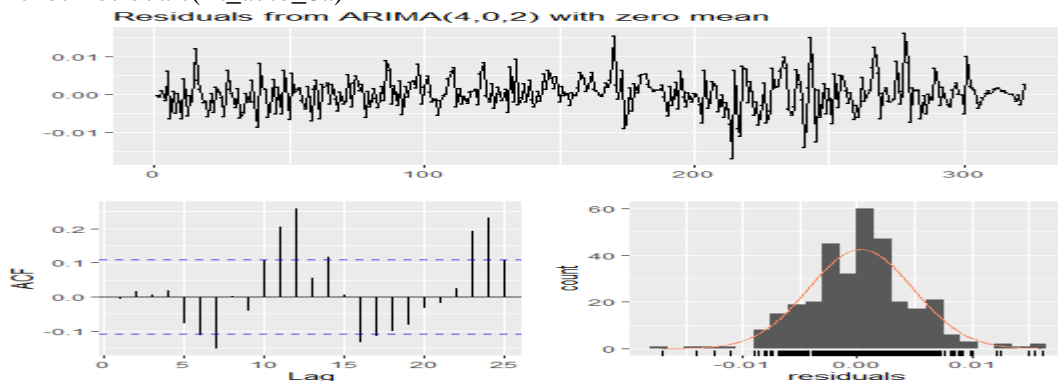
Coefficients:

	ar1	ar2	ar3	ar4	ma1	ma2
	-0.3853	0.7049	0.5661	-0.2001	1.4618	0.8427
s.e.	0.0948	0.0835	0.0805	0.0651	0.0794	0.0583

sigma² estimated as 2.116e-05: log likelihood=1277.78

AIC=-2541.55 AICc=-2541.2 BIC=-2515.13

```
> checkresiduals(fit_auto_8a)
```

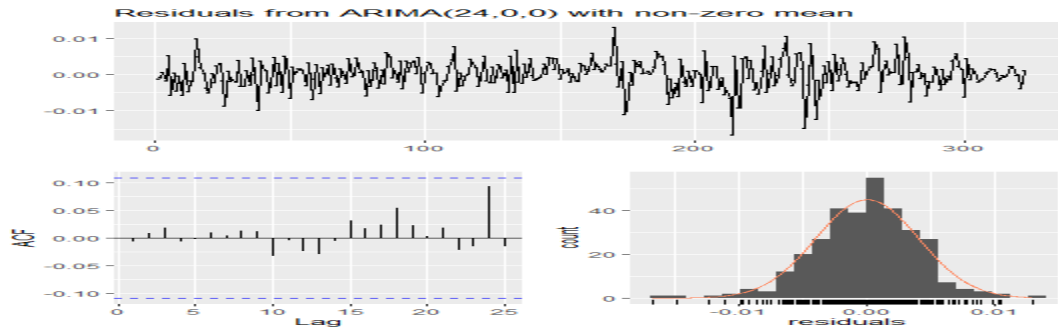


Conclusion: A lot of significant autocorrelations in data. There is a repeated wave pattern.

b

```
fit_arima_8a <- arima(x_sf, order=c(24,0,0))
```

```
> checkresiduals(fit_arima_8a)
```



Note: No visible autocorrelation, but a lot of parameters in the model (25). In general, large models are not very efficient. A SARIMA model may work better.

c. Forecast SF_c for the period 2017:1 to 2019:9.

```
sf_ar24_coeff <- fit_arima_8a$coef
```

```
lag <- 24
```

```
x0_lag <- matrix(1,(T-lag),1)
```

```
xx_sf_lag <- cbind(sf_1, sf_2, sf_3, x0_lag)
```

```
sf_lag <- matrix(0, (T-lag),lag)
```

```
j <- 1
```

```
while (j <= lag) { # loop to create 24 lags
```

```
sf_lag[j] <- x_sf1[j:(T-lag+j-1)]
```

```
j <- j + 1
```

```
}
```

```
y_f0_arlag <- xx_sf_lag[302:nrow(xx_sf_lag),]%% sf_ar24_coeff
```

```
err_f0_arlag <- y[302:nrow(xx_sf_lag)] - y_f0_arlag
```

```
mse_arlag <- sum(err_f0_arlag^2)/length(y_f0_arlag)
```

```
> mse_arlag
```

```
[1] 0.0001270861 => Higher MSE
```

d.

```
> dm.test(err_f0_arlag,err_f0_ar3, power=2)
```

Diebold-Mariano Test

```
data: err_f0_arlagerr_f0_ar3
```

```
DM = 2.6119, Forecast horizon = 1, Loss function power = 2, p-value = 0.01393
```

```
alternative hypothesis: two.sided
```

Conclusion: AR(3) has a better MSE.